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Convergence in Distribution of Multiply-indexed Arrays, with Applications in MANOVA

Zbieżność według rozkładu wielowskaźnikowych tablic z zastosowaniami w MANOVA

Сходимость по распределению мультииндексных таблиц с приложениями в MANOVA

1. Introduction. The importance of convergence in distribution in statistical inference arises as follows. The data  $y_1, \dots, y_n$  arising from  $n$  performances of a given random process  $\epsilon$  is used to calculate various quantities of interest,  $W_n = (W_{1n}, \dots, W_{in})'$  say, which are then used to construct significance, confidence intervals etc. These require the evaluation of probabilities of the form  $P(W_n \in A)$ , for given sets  $A \in R^I$ . If the distribution of  $W_n$  is intractable, an approximation to  $P(W_n \in A)$  is available when the sample size  $n$  is large in the case when the sequence  $\{W_n\}$  converges in distribution to a variate  $W$  with known distribution for then ([2], Theorem 2.1)

$$\lim_{n \rightarrow \infty} P(W_n \in A) = P(W \in A)$$

for all sets  $A$  of practical interest.

Consider now the situation in e.g. MANOVA. There are now several ( $k$  say) independent random processes  $\epsilon_1, \dots, \epsilon_k$ , the data arises from  $n_i$  performances of  $\epsilon_i$ ,  $i = 1, \dots, k$ , and leads to quantities of interest of the form  $W_{n_1, \dots, n_k}$ . An approximation to  $P(W_{n_1, \dots, n_k} \in A)$  when all the sample sizes  $n_1, \dots, n_k$  are large may then be important in practice for similar reasons. Such approximations are provided by the type of convergence in distribution of multiply-indexed arrays  $\{W_{n_1, \dots, n_k}\}$  of random vectors that is defined below.

2. Multiply-indexed arrays. We discuss in some detail only the case  $k = 2$ , since the treatment when  $k > 2$  presents no additional difficulties.

**2. 1 Definition and notation.** We call a set of real numbers  $\{a_{n_1, n_2}; n_1 \geq 1, n_2 \geq 1\}$  a doubly-indexed array. It may be conveniently pictured in table form –

$n_2$	1	2	3	...
1	$a_{11}$	$a_{12}$	$a_{13}$	...
2	$a_{21}$	$a_{22}$	$a_{23}$	...
3	$a_{31}$	$a_{32}$	$a_{33}$	...
⋮	⋮	⋮	⋮	⋮

In view of later use in MANOVA, it will be convenient to write  $N = \text{diag}(n_1, n_2)$ ,  $a_{n_1, n_2} = \{a_N\}$ , and the array  $a_N$ .

Further, if  $N_1 = \text{diag}(n_{11}, n_{12})$ ,  $N_2 = \text{diag}(n_{21}, n_{22})$ , we shall write  $N_1 > N_2$  if  $n_{11} > n_{21}$  and  $n_{12} > n_{22}$ , with a similar meaning for  $N_1 \geq N_2$ .

Finally by ' $N$  is arbitrarily large' we shall mean that  $n_1$  and  $n_2$  are both arbitrarily large.

**2.2 Limit points.** We say that  $\alpha$  (finite) is a limit point of  $\{a_N\}$  if for arbitrary  $\epsilon > 0$  and  $n > 0$ ,  $\exists N \geq nI$  such that  $|a_N - \alpha| < \epsilon$ .

There is the usual extension to infinite limit points. (There will be similar extension below, which in general will not be mentioned explicitly.)

We show now, by a standard argument, that every array  $\{a_N\}$  has a limit point.

In the case when  $\{a_N\}$  is bounded,  $a_N \in J_0 = [a, b] \forall N$  say, construct a sequence  $\{J_n\}$  of closed intervals by repeated subdivision of  $J_0$ , viz., for  $n = 1, 2, \dots$ ,  $J_n$  is the left half of  $J_{n-1}$  if this half contains terms  $a_N$  with  $N$  arbitrarily large, and otherwise  $J_n$

is the right half. Then  $\{J_n\}$  defines the point  $\alpha = \bigcap_1 J_n$ . This point  $\alpha$  is a limit point of

$\{a_N\}$ , since  $\forall n J_n$  contains  $\alpha$  and terms  $a_N$  with  $N$  arbitrarily large. Moreover,  $\alpha = \liminf a_N$ , since for arbitrary  $\epsilon > 0$ ,  $\exists n_0$  such that  $a_N > \alpha - \epsilon \forall N \geq n_0I$ .

If  $\{a_N\}$  is not bounded, a similar argument shows that  $a_N$  has a limit point, which now may be infinite.

**2.3 Subarrays.** Let  $S$  be a subset of diagonal matrices  $N$  that contains matrices that are arbitrarily large. We call  $\{a_N, N \in S\}$  a subarray of  $\{a_N\}$ .

Limit points of subarrays are defined in the obvious way, and it follows, as in 2.2, that every subarray has a limit point.

**2.4 Convergence.** We say that  $\{a_N\}$  converges to  $\alpha$  (finite), and write  $\lim_{N \rightarrow \infty} a_N = \alpha$ ,

if for arbitrary  $\epsilon > 0$ ,  $\exists n_0$  such that  $|a_N - \alpha| < \epsilon \forall N \geq n_0I$ .

Similarly, we say that the subarray  $\{a_N, N \in S\}$  converges to  $\alpha$  (finite) if for arbitrary  $\epsilon > 0$ ,  $\exists n_0$  such that  $|a_N - \alpha| < \epsilon \forall N \in S$  such that  $N \geq n_0I$ .

The usual results then follow. As an example, we prove that if  $\alpha$  is a limit point of  $\{a_N\}$  then there exists a subarray that converges to  $\alpha$ .

In the case when  $\alpha$  is finite, let  $\{\epsilon_i\}$  be a null sequence of positive terms, and construct a set  $S = \{N_i\}$  as follows.

Choose  $N_1 > I$  such that  $|a_{N_1} - \alpha| < \epsilon_1$ , then successively choose  $N_{i+1} > N_i$  such that  $|a_{N_{i+1}} - \alpha| < \epsilon_{i+1}$ ,  $i = 1, 2, \dots$ , (such  $N_i$  always exist, from 2.2).

For given  $\epsilon > 0$ ,  $\exists j$  such that  $\epsilon_i < \epsilon \forall i \geq j$ . Then  $|a_{N_j} - \alpha| < \epsilon \forall i \geq j$ .

Writing  $n_0 = \min(n_{j_1}, n_{j_2})$ , where  $N_j = \text{diag}(n_{j_1}, n_{j_2})$ , then  $S \cap \{N, N \geq n_0 I\} = \{N_i, i \geq j\}$ , whence  $|a_N - \alpha| < \epsilon \forall N \in S$  such that  $N \geq n_0 I$  and the subarray  $\{a_N, N \in S\}$  converges to  $\alpha$ .

There is a similar result for limit points of subarrays.

We mention one further result, viz., that if  $\{a_N\}$  converges to  $\alpha$ , then every subarray of  $\{a_N\}$  converges to  $\alpha$ . And there is the corresponding result for convergent subarrays.

**2.5 Lim inf and lim sup.** The following treatment is parallel to Feller's treatment of lim inf and lim sup ([4], IV. 1), and uses his  $\cap, \cup$  notation.

We first introduce a sequential ordering of the terms of  $\{a_N\}$  with  $N \geq nI$ , viz.

$$a_{nn}, a_{n+1n}, a_{nn+1}, a_{n+2n}, a_{n+1n+1}, a_{nn+2}, a_{n+3n}, \dots$$

Next, consider the sequence  $\{w_n\}$ , where

$$w_n = a_{nn} \cap a_{n+1n} \cap a_{nn+1} \cap a_{n+2n} \cap \dots = \bigcap_{N \geq nI} a_N$$

Clearly  $w_n \uparrow$ , whence  $\{w_n\}$  convergence to a limit,  $\alpha$  say. Thus, in the case when  $\alpha$  is finite, for arbitrary

$$\epsilon > 0, \exists n_0 \text{ such that } \alpha - \epsilon < w_n \leq \alpha \forall n \geq n_0 \tag{1}$$

We now show that for arbitrary  $\epsilon > 0$ ,

$$\exists n_1 \text{ such that } a_N > \alpha - \epsilon \forall N \geq n_1 I, \text{ and} \tag{i}$$

$$\exists N \text{ arbitrarily large such that } a_N < \alpha + \epsilon, \tag{ii}$$

from which it follows that  $\alpha = \lim \inf a_N$ . Firstly, since by definition  $w_{n_0} \leq a_N \forall N \geq n_0 I$ , then, from (1), (i) holds with  $n_1 = n_0$ . Next, suppose that (ii) does not hold. Then  $\exists \epsilon_1 > 0$  and  $n_2$  such that  $a_N \geq \alpha + \epsilon_1 \forall N \geq n_2 I$ . But then  $w_n \geq w_{n_2} \geq \alpha + \epsilon_1 \forall n \geq n_2$ , which contradicts (1).

There is a similar treatment for  $\lim \sup a_N$ .

**2.6 Fatou's lemma and the dominated convergence theorem.** We consider now an array  $\{f_N(x)\}$  of functions  $f: R^I \rightarrow R^1$ . Then, from 2.5, for each  $x$

$$w_n(x) = \bigcap_{N \geq nI} f_N(x)$$

defines an increasing sequence  $\{w_n(x)\}$  that converges to  $\lim \inf f_N(x)$ .

Fatou's lemma. ([4], IV. 2) Suppose that  $\{f_N(x)\}$  is an array of non-negative functions, and that  $F(x)$  is a distribution function (d.f.).

If  $f_N$  is integrable for all  $N$ , i.e. if

$$E[f_N] = \int_{\mathbb{R}^2} f_N(x) dF(x) < \infty \quad \forall N$$

then

$$E[\liminf f_N] \leq \liminf E[f_N].$$

**Proof.** Define a sequence of functions  $\{f_n\}$  as follows. For each  $n \geq 1$ , choose  $N_n$  such that  $N_n \geq nl$ , and define

$$f_n = f_{N_n} \quad (2)$$

By definition of  $w_n$ ,  $w_n \leq f_n \quad \forall n$ , whence  $E(w_n) \leq E(f_n) \quad \forall n$ , and so

$$\liminf E(w_n) \leq \liminf E(f_n). \quad (3)$$

Since  $w_n \uparrow$ ,  $\lim w_n = \liminf f_N$ , and  $w_n$  is integrable for all  $n$ , then, by the monotone convergence theorem ([4], IV. 2),  $\{E(w_n)\}$  converges, and  $\lim E(w_n) = E(\lim w_n)$ .

It follows then, using (3), that

$$E(\liminf f_n) \leq \liminf E(f_n). \quad (4)$$

Since (4) holds for all sequences  $\{f_n\}$  satisfying (2), it is enough to show that there exists such a sequence for which  $\liminf E(f_n) = \liminf E(f_N)$ . To show this, consider the array  $\{E(f_N)\}$ , and write  $\alpha = \liminf E(f_N)$ . By 2.4, there exists a subarray  $\{E(f_N), N \in S\}$  that converges to  $\alpha$ . Consider now the corresponding subarray  $\{f_N, N \in S\}$ , and construct from it a sequence  $\{f_n\}$  as follows. Choose any element  $N_1$  of  $S$  and define  $f_1 = f_{N_1}$ ; then for  $n = 1, 2, \dots$ , choose an element  $N_{n+1}$  of  $S$  such that  $N_{n+1} > N_n$  and define  $f_{n+1} = f_{N_{n+1}}$ . Then

(i)  $N_n \geq nl \quad \forall n$ , so that  $\{f_n\}$  satisfies (2).

(ii) Since  $\{E(f_n)\} = \{E(f_N), N \in S_1\}$ , where  $S_1 = \{N_i\} \subset S$ , and, by 2.4, the subarray  $\{E(f_N), N \in S_1\}$  converges to  $\alpha$ , then  $\lim_{n \rightarrow \infty} E(f_n) = \alpha$  and  $\liminf E(f_n) = \liminf E(f_N)$ , as required.

The following theorem then follows from Fatou's lemma in the standard way (see e.g. [4], IV. 2).

**Dominated convergence theorem.** If  $\{f_N(x)\}$  is an array such that  $f_N$  is integrable  $\forall N$ , and that  $\lim_{N \rightarrow \infty} f_N(x) = f(x)$  pointwise, and that there exists an integrable function  $u$

such that  $|f_N(x)| < u(x) \quad \forall x$ , then

$$\lim_{N \rightarrow \infty} E(f_N) = E(f).$$

**2.7 Helly's theorem.** Helly's theorem ([4], VIII. 6) may be generalized to arrays  $\{F_N(x)\}$  of d.f. The proof is essentially the same as the proof in Feller, and depends on the following lemma.

**Lemma.** If  $\{f_N(x)\}$  is a given array of bounded functions ( $R^k \rightarrow R^l$ ) and  $\{a_i\}$  is a given sequence of points in  $R^k$ , then there exists a subarray  $\{f_N, N \in S\}$  that converges at all points  $a_i$ .

**Proof.** (c.f. [4], VIII. 6). By 2.2, the bounded array  $\{f_N(a_i)\}$  has a limit point, and hence contains a convergent subarray  $\{f_N(a_1), N \in S_1\}$ . Proceeding in this way, the bounded subarray  $\{f_N(a_2), N \in S_1\}$  has a limit point, and hence contains a convergent subarray  $\{f_N(a_2), N \in S_2\}$ . Continuing this procedure, we generate a sequence of sets  $S_1 \supset S_2 \supset \dots \supset S_n \supset \dots$  such that for each  $i = 1, 2, \dots, \{f_N(a_i), N \in S_i\}$  is a convergent subarray.

For each  $n \geq 1$  we now choose an element  $N_n \in S_n$  such that  $N_n \geq nl$ , and define  $S = \{N_n\}$ . Then the subarray  $f_N, N \in S$  has the required property. For consider  $\{f_N(a_i), N \in S\}$ . Since  $N_n \in S_n \subset S_i \forall n \geq i$ , then apart from a finite number of terms,  $\{f_N(a_i), N \in S\}$  is a subarray of  $\{f_N(a_i), N \in S_i\}$  which we know converges.

Thus  $\{f_N(a_i), N \in S\}$  converges for  $i = 1, 2, \dots$

The generalizations of these results when  $k > 2$  are now used to develop a theory of convergence in distribution for multiply-indexed arrays.

**3. Convergence in distribution for multiply-indexed arrays.** Let  $\{W_N\}$  be a  $k$ -fold multiply-indexed array of  $l \times l$  vector variates and  $W$  an  $l \times l$  vector variate. We denote the corresponding d.f. by  $\{F_N(x)\}$  and  $F(x)$ , the corre-characteristic functions (c.f.) by  $\{\zeta_N(t)\}$  and  $\zeta(t)$ , and write

$$E_N(f) = \int_{R^l} f(x) dF_N(x)$$

and

$$E(f) = \int_{R^l} f(x) dF(x).$$

**Definition.** We say that  $\{W_N\}$  converges in distribution to  $W$ , and write  $W_N \xrightarrow{D} W$ , iff  $\lim_{N \rightarrow \infty} F_N(x) = F(x) \forall$  continuity points  $x$  of  $F$ .

**Theorem 1.**  $W_N \xrightarrow{D} W$  if and only if either

(i)  $\lim_{N \rightarrow \infty} P(W_N \in I) = P(W \in I)$  for all bounded open 'rectangles'  $I$  such that  $P(W \in \partial I) = 0$ ,

or (ii)  $\lim_{N \rightarrow \infty} P(W_N \in A) = P(W \in A)$  for all Borel sets  $A$  such that  $P(W \in \partial A) = 0$ ,

or (iii)  $\lim_{N \rightarrow \infty} E_N(f) = E(f)$  for all bounded and continuous functions  $f: R^l \rightarrow R^1$ ,

or (iv)  $\lim_{N \rightarrow \infty} \zeta_N(t) = \zeta(t) \forall t$ .

Moreover, if  $W_N \xrightarrow{D} W$ , then the convergence in (iv) is uniform for all  $t$  in any bounded domain of  $R^l$ .

**Proof.** The proof of (ii) and (iii) depends only on the content of 2.1 – 2.4 and follows step for step the corresponding proof of Bilingsley ([2], §2). The fact that (i)  $\Rightarrow$  (iii) similarly follows the proof of the theorem in [4], VIII. 1. The fourth part (a continuity theorem for c.f.) depends also on 2.6 – 2.7, and can be proved in the same way as the corresponding 'ordinary' theorem, as e.g. in [4], XV. 3 or in [3], Chapter 11.

In the case when  $\underline{W}$  has a singular distribution concentrated at the single point  $\underline{\alpha}$ , we say that  $\{\underline{W}_N\}$  converges in probability to  $\underline{\alpha}$ , and we shall write  $\underline{W}_N \xrightarrow{D} \underline{\alpha}$ , as well as the standard  $\underline{W}_N \xrightarrow{P} \underline{\alpha}$ .

All the standard results for convergence in distribution of sequences of vector variates have their obvious counterparts in the theory of convergence in distribution of multiply-indexed arrays. We recall in particular two results. The first states that, if  $\underline{W}_{1N} \xrightarrow{D} \underline{W}_1$  and  $\underline{W}_{2N} \xrightarrow{D} \underline{\alpha}_2$ , then writing  $\underline{W}_N = (\underline{W}'_{1N}, \underline{W}'_{2N})'$ ,  $\underline{W}_N \xrightarrow{D} \underline{W} = (\underline{W}'_1, \underline{W}'_2)'$ , where  $P(\underline{W}_2 = \underline{\alpha}_2) = 1$ , i.e. the limiting joint distribution is singular, and concentrated on the hyperplane  $\underline{W}_2 = \underline{\alpha}_2$ . In such a case, we shall write

$$\begin{pmatrix} \underline{W}_{1N} \\ \underline{W}_{2N} \end{pmatrix} \xrightarrow{D} \begin{pmatrix} \underline{W}_1 \\ \underline{\alpha}_2 \end{pmatrix}$$

The second result, which has widespread application, we state as a theorem.

**Theorem 2.**

$$\underline{W}_N \xrightarrow{D} \underline{W} = \phi(\underline{W}_N) \xrightarrow{D} \phi(\underline{W})$$

for every Borel-measurable function  $\phi: R^q \rightarrow R^q$  such that  $P(\underline{W} \in D_\phi) = 0$ , where

$$D_\phi = \{x; \phi(x) \text{ is discontinuous}\}.$$

The proof again follows step for step the corresponding proof in Billingsley ([2], Corollary 3 of theorem 3.3).

**4. Some asymptotic results in MANOVA.** As an application of §3, we now derive some asymptotic results in MANOVA, on the assumption of a common non-singular covariance matrix  $\Sigma$ .

**4.1 MANOVA notation.** We suppose that the data is obtained from  $n_i$  performances of the random process  $\mathcal{E}_i$ ,  $i = 1, \dots, k$ , where  $\mathcal{E}_1, \dots, \mathcal{E}_k$  are independent processes. For  $\mathcal{E}_i$ , we denote the  $p \times 1$  variate by  $\underline{y}_i$  and its mean by  $\underline{\mu}_i$ . We denote the corresponding  $n_i \times p$  data matrix by  $Y_i$ , and the sample mean and covariance matrix by  $\bar{y}_i$  and  $S_{(i)}$ . We write

$$\begin{aligned} \sum_1^k n_i &= n, & N &= \text{diag}(n_1, \dots, n_k), \\ M_{k \times p} &= \begin{pmatrix} \underline{\mu}'_1 \\ \vdots \\ \underline{\mu}'_k \end{pmatrix} = (\mu_{ij}), & \bar{Y}_{N \times p} &= \begin{pmatrix} \underline{y}'_1 \\ \vdots \\ \underline{y}'_k \end{pmatrix} = (y_{ij}), \\ y_{i \times p} &= \begin{pmatrix} \underline{y}'_{i1} \\ \vdots \\ \underline{y}'_{in_i} \end{pmatrix}, & Y_{n \times p} &= \begin{pmatrix} Y_1 \\ \vdots \\ Y_k \end{pmatrix}, \end{aligned}$$

and  $S_{p \times p} = \frac{1}{n-k} \sum_i (n_i - 1) S_{(i)} = \frac{1}{n-k} \sum_{i=1}^k \sum_{j=1}^{n_i} (y_{ij} - \bar{y}_i) (y_{ij} - \bar{y}_i)'$ .

The above assumptions, which we shall call the model  $G$ , can be summed up as follows:

$G$ : The rows of  $Y$  are independent vector variates, and  $E(Y) = XM, \text{Var}(Y) = \Sigma \otimes I_n$ , where

$$X = \begin{pmatrix} I_1 & 0 & \dots & 0 \\ 0 & I_2 & \dots & 0 \\ \vdots & \vdots & \dots & \vdots \\ 0 & 0 & \dots & I_k \end{pmatrix} \quad (5)$$

$n \times k$

Note that

$$X'X = N, X'Y = N\bar{Y}_N \text{ and } r(X) = k \quad (6)$$

and further, that each column of  $E(Y) \subset \mathcal{R}(X) \subset R^n$ , that  $P = X(X'X)^{-1} X' = XN^{-1} X'$  is the orthogonal projector matrix (o.p. matrix) onto  $\mathcal{R}(X)$ , and that  $(n-k)S = Y'(I-P)Y$ .

We now consider the usual kind of MANOVA hypothesis  $H$ , viz.

$$H: M = X_1 B_1 \text{ where } X_1 \text{ is a known } k \times r \text{ matrix of rank } r. \quad (7)$$

When  $H$  is true,  $E(Y) = XX_1 B_1 = X_0 B_1$  where  $X_0 = XX_1$  has rank  $r$ , each column of  $E(Y) \subset \mathcal{R}(X_0) \subset \mathcal{R}(X)$ , the o.p. matrix onto  $\mathcal{R}(X_0)$  is

$$P_0 = XX_1 (X_1' N X_1)^{-1} X_1' X'$$

and

$$S_0 = \frac{1}{n-r} Y'(I-P_0)Y$$

is an unbiased estimate of  $\Sigma$ .

The MANOVA table for testing  $H$  is then

Source	SSP	DF	MSSP
$H$ vs. $G$	$Y'(P-P_0)Y = S_1$	$k-r$	$S_1$
Within class	$Y'(I-P)Y = S$	$n-k$	$S$
Total	$Y'(I-P_0)Y = S_0$	$n-r$	$S_0$

4.2 A central limit theorem.

Theorem 3. On  $G, N^{1/2}(\bar{Y}_N - M) \xrightarrow{D} W \sim N(0, \Sigma \otimes I_k)$ .

Proof. Writing

$$T_{k \times p} = \begin{pmatrix} t'_1 \\ \vdots \\ t'_k \end{pmatrix}$$

the c.f.  $\zeta_N(T)$  of  $W_N = N^{1/2}(\bar{Y}_N - M)$  is

$$\begin{aligned} \zeta_N(T) &= E[\exp(i \text{Tr}(T' W_N))] = E\left[\prod_{j=1}^k (\exp(i t'_j(\bar{y}_j - \mu_j) \sqrt{n_j}))\right] = \\ &= \prod_{j=1}^k E[\exp(i t'_j(\bar{y}_j - \mu_j) \sqrt{n_j})] = \prod_{j=1}^k \phi_{n_j}^{(j)}(t_j) \end{aligned}$$

since  $\bar{y}_1, \dots, \bar{y}_k$  are independent where

$$\phi_{n_j}^{(j)}(t_j) = E[\exp(i t'_j(\bar{y}_j - \mu_j) \sqrt{n_j})], \quad j = 1, \dots, k.$$

But it is known from the ordinary multivariate central limit theorem that, for given  $t_j$ ,

$$\lim_{n_j \rightarrow \infty} \phi_{n_j}^{(j)}(t_j) = \exp(-0.5 t'_j \Sigma_j t_j), \quad j = 1, \dots, k$$

Thus, for given  $T$ ,

$$\lim_{N \rightarrow \infty} \zeta_N(T) = \prod_{j=1}^k \exp(-0.5 t'_j \Sigma_j t_j) = \exp(-0.5 \text{Tr}(T \Sigma T')) = E[\exp(i \text{Tr}(T' W))],$$

and the theorem follows from theorem 1.

**4.3 The asymptotic distribution of  $S_1$ .** Suppose that  $A$  is a  $p \times p$  symmetric matrix. By  $\underline{A}$  we shall mean the  $p(p+1)/2 \times 1$  vector

$$A = (a_{11}, \dots, a_{1p}, a_{22}, \dots, a_{2p}, \dots, a_{pp}).$$

**Theorem 4.** When  $H$  is true,  $S_1 \xrightarrow{D} \underline{V}$ , where  $V = U'U$  and  $U \sim N(0, \Sigma \otimes I_{k-r})$ .  
 $(k-r) \times p$

**Proof.** When  $H$  is true, the columns of  $XM$  lie in  $\mathcal{R}(X_0) \subset \mathcal{R}(X)$ , so that  $(P - P_0)XM = 0$ . Thus

$$S_1 = Y'(P - P_0)Y = (Y - XM)'(P - P_0)(Y - XM) = W'_N(I - P_N)W_N,$$

where, from (6) and the definitions of  $P$  and  $P_0$ ,

$$P_N = N^{1/2} X_1 (X_1' N X_1)^{-1} X_1' N^{1/2},$$

$$k \times k$$

the o.p. matrix onto the  $r$ -dimensional subspace  $\Omega_N$  of  $R^k$ , where  $\Omega_N = \mathcal{R}(N^{1/2} X_1)$ .

Now let  $H_N$  be a matrix whose columns are an orthonormal basis of  $\Omega_N$ . Then  
 $k \times (k-r)$

$$H'_N H_N = I_{k-r}, \quad H_N H'_N = I - P_N, \text{ and}$$



$$S_1 = U'_N U_N \tag{8}$$

where  $U_N = H'_N W_N$ .

We now show that  $U_N \xrightarrow{D} U \sim N(0, \Sigma \otimes I_{k-r})$ , from which the theorem follows by a simple application of theorem 2. From Theorems 3 and 1

$$\zeta_N(T) = \exp(-0.5 Tr(T \Sigma T')) + f_N(T) \tag{9}$$

where  $\lim_{N \rightarrow \infty} f_N(T) = 0$  uniformly for  $T$  in any bounded domain  $A \subset R^{kp}$ .

Consider now the c.f.  $\phi_N(T_1)$  of  $U_N$ , where  $T_1$  is  $(k-r) \times p$ .

$$\begin{aligned} \phi_N(T_1) &= E[\exp(i Tr(T'_1 U_N))] = \zeta_N(H_N T_1), \text{ using (8)} \\ &= \exp(-0.5 Tr(H_N T_1 \Sigma T'_1 H'_N)) + f_N(H_N T_1) = \exp(-0.5 Tr(T_1 \Sigma T'_1)) + f_N(H_N T_1) \end{aligned}$$

since  $H'_N H_N = I_{k-r}$ .

For fixed  $T_1$ , choose in (9)  $A = \{T; Tr(T' T) \leq Tr(T'_1 T_1)\}$ . Since  $Tr((H_N T_1)'(H_N T_1)) = Tr(T'_1 T_1) \forall N$ , then, from (9)  $\lim_{N \rightarrow \infty} f_N(H_N T_1) = 0$ , and the result follows by an

application of Theorem 1.

**Theorem 5.** On  $G, \underline{\Sigma} \xrightarrow{D} \underline{\Sigma}$ .

**Proof.** We write  $\nu_j = n_j - l, j = l, \dots, k$  and  $\nu = n - k$ , so that

$$S = \sum_{j=1}^k (\nu_j/\nu) S_{(j)}.$$

It is well-known that for each  $j \underline{\Sigma}_{(j)} \xrightarrow{D} \underline{\Sigma}$  as  $n_j \rightarrow \infty$ . Thus, writing

$$\phi_{n_j}^{(j)}(t_j) = E[\exp(i t'_j \underline{\Sigma}_{(j)})] = \exp(i t'_j \underline{\Sigma}) + f_{n_j}^{(j)}(t_j),$$

then  $\lim_{n_j \rightarrow \infty} f_{n_j}^{(j)}(t_j) = 0$  uniformly for  $t_j$  in any bounded domain.

Now write  $g_{n_j}^{(j)}(t) = \exp(-i t'_j \underline{\Sigma}) f_{n_j}^{(j)}(t), j = l, \dots, k$ .

Then

$$\phi_{n_j}^{(j)}(t) = (1 + g_{n_j}^{(j)}(t)) \exp(i t'_j \underline{\Sigma})$$

and, since  $|\exp(i t'_j \underline{\Sigma})| = 1$ , then, given  $\epsilon_1 > 0, k > 0, \exists n_{0j}$  such that  $|g_{n_j}^{(j)}(t)| < \epsilon_1 \forall n_j \geq n_{0j}$  and

$$\forall t \in A = \{t; t'_j t \leq k\}. \tag{10}$$

Consider now the c.f.  $\zeta_N(t)$  of  $S$ , viz.

$$\zeta_N(t) = E\left[\prod_{j=1}^k \exp(i t' (v_j/\nu) S_{nj})\right] = \prod_{j=1}^k \phi_{nj}^{(j)}(v_j t/\nu),$$

since  $S_1, \dots, S_k$  are independent. Thus (all logs being principal-valued)

$$\log \zeta_N(t) = (i t' \Sigma) \sum_j (v_j/\nu) + \sum_j \log(1 + g_{nj}^{(j)}(v_j t/\nu)) + (2C_N \pi) i,$$

where  $C_N$  is integer depending on  $N$ .

Since  $\Sigma(v_j/\nu) = 1$ , the theorem will follow by showing that, for fixed  $t$ ,

$$\lim_{N \rightarrow \infty} \sum_j \log(1 + g_{nj}^{(j)}(v_j t/\nu)) = 0.$$

Using the fact that

$$|\log(1+z)| < 2|z| \text{ if } |z| < 0.5,$$

then

$$|\sum_j \log(1 + g_{nj}^{(j)}(v_j t/\nu))| < 2 \sum_j |g_{nj}^{(j)}(v_j t/\nu)|$$

provided that

$$|g_{nj}^{(j)}(v_j t/\nu)| < 0.5, \quad j = 1, \dots, k.$$

For arbitrary  $\epsilon > 0$ , choose now  $\epsilon_1 = \epsilon/2k$  and  $k = \lfloor t' t \rfloor$  in (10), and write  $n_0 = \max(n_{01}, \dots, n_{0k})$ . Since  $v_j t/\nu \in A \forall N$ , it then follows from (10) that

$$2 \sum_j |g_{nj}^{(j)}(v_j t/\nu)| < \epsilon \forall N \geq n_0$$

and hence that

$$\lim_{N \rightarrow \infty} \sum_j \log(1 + g_{nj}^{(j)}(v_j t/\nu)) = 0.$$

**4.4 The eigen-values of  $S_1 S^{-1}$ .** We now consider the asymptotic distribution of the

e. values of  $S_1 S^{-1}$  when  $H$  is true. Since Theorem 5  $\Rightarrow |S| \xrightarrow{D} |\Sigma| > 0$ , it follows that the possible lack of definition of  $S^{-1}$  has no effect on the asymptotic distribution. Furthermore, since

$$r(S_1) \leq \rho = \min(p, k-r) \forall N,$$

with equality almost always when  $n$  is large, only the  $\rho$  largest e. values  $\lambda_1 \geq \lambda_2 \geq \dots \geq \lambda_\rho$  are of interest.

**Theorem 6.** When  $H$  is true,  $\underline{\lambda}_N \xrightarrow{D} L$ , where  $\underline{\lambda}_N = (\lambda_1, \dots, \lambda_\rho)'$ ,  $L = (L_1, \dots, L_\rho)'$ ,  $L_1 \geq L_2 \geq \dots \geq L_\rho$  are the largest e. values of  $Z'Z$ , and  $Z \sim N(0, I_{p(k-r)})$ .

**Proof.** From theorem 4 and 5

$$\begin{pmatrix} \underline{S}_1 \\ S \end{pmatrix} \xrightarrow{D} \begin{pmatrix} Y \\ \Sigma \end{pmatrix} \text{ when } H \text{ is true.}$$

Since  $\underline{\lambda}_N = \phi(\underline{S}_1, \underline{S})$ , where  $\phi$  is Borel-measurable and continuous when  $\underline{S} = \underline{\Sigma}$ , it follows from Theorem 2 that  $\underline{\lambda}_N \xrightarrow{D} \phi(\underline{V}, \underline{\Sigma})$ , i.e. the vector of the  $\rho$  largest e. values of  $V\Sigma^{-1} = U'U\Sigma^{-1}$ , where from Theorem 1.4,

$$U \sim N(0, \Sigma \otimes I_{k-r}).$$

Write now  $\Sigma^{-1} = A^2$ , where  $A$  is symmetric. Since  $V\Sigma^{-1} = A^{-1}(AU'UA)A = A^{-1}(Z'Z)A$ , where  $Z = UA \sim N(0, I_{p(k-r)})$ , then  $V\Sigma^{-1}$  and  $Z'Z$  have the same e. values, and the result follows.

This theorem allows us to write down the asymptotic distribution when  $H$  is true of some statistics commonly used in practice for testing  $H$ , viz. Hotelling's  $T_0^2$ , Pillai's  $V^{(p)}$ , and the statistic  $U$ , which is essentially the Normal theory likelihood-ratio statistic, where

$$T_0^2 = Tr(S_1 S^{-1}) = \sum_i \lambda_i, \quad V^{(p)} = \frac{n-k}{n-r} Tr(S_1 S_0^{-1}) = \sum_i \frac{\lambda_i}{1 + (\lambda_i/n - k)}$$

$$\text{and } U^{-1} = \prod_i \left(1 + \frac{\lambda_i}{n-k}\right)$$

see e.g. [1], Ch. 8).

Theorem 7. When  $H$  is true,  $T_0^2$ ,  $V^{(p)}$  and  $(n-k)(U^{-1} - I)$  each converges in distribution to

$$Tr(Z'Z) \sim \chi_{p(k-r)}^2.$$

Proof. It follows immediately from Theorem 2 that

$$T_0^2 = \sum_i \lambda_i \xrightarrow{D} \sum_i L_i = Tr(Z'Z) \sim \chi_{p(k-r)}^2$$

and similarly that

$$V^{(p)} \xrightarrow{D} \sum_i L_i / (1 + 0.L_i) = Tr(Z'Z).$$

Finally,

$$(n-k)(U^{-1} - I) = \sum_i \lambda_i + (n-k)^{-1} \sum_{i \neq j} \lambda_i \lambda_j + \dots + \frac{1}{(n-k)^{p-1}} \prod_i \lambda_i \\ \xrightarrow{D} \sum_i L_i + 0. \sum_{i \neq j} L_i L_j + \dots + 0. \prod L_i = Tr(Z'Z).$$

(It can also be shown somewhat similarly that

$$-n \log U \xrightarrow{D} \chi_{p(k-r)}^2.)$$

4.5 Estimation of  $B_1$ . If  $H$  is not rejected, the estimation of  $B_1$  will often be of importance. We consider the asymptotic distribution of  $\hat{B}_N$  when  $H$  is true, where  $\hat{B}_N = (X_0'X_0)^{-1} X_0'Y$  is the matrix of minimum variance unbiased linear estimates of  $B_1$  when  $H$  is true. Since  $E(\hat{B}_N) = (X_0'X_0)^{-1} X_0'X_0 B_1 = B_1$  and

$$\text{Var}(\hat{B}_N) = (I_p \otimes (X'_0 X_0)^{-1} X'_0) (\Sigma \otimes I_n) (I_p \otimes X_0 (X'_0 X_0)^{-1}) = \Sigma \otimes (X'_1 N X_1)^{-1},$$

it might be expected that  $\hat{B}_N$  is asymptotically  $N(B_1, \Sigma \otimes (X'_1 N X_1)^{-1})$ , in the sense that, if  $C_N$  and  $A$  are respectively  $r \times r$  and  $p \times p$  symmetric matrices such that

$$C_N^2 = X'_1 N X_1, \quad A^2 = \Sigma^{-1} \quad (11)$$

then  $C_N(\hat{B}_N - B_1)A \xrightarrow{D} Z_1 \sim N(0, I_{pr})$ .

To prove this, note first that when  $H$  is true  $E(Y) = XM = X_0 B_1$ , so that, using (6) and the notation of Theorem 3,

$$\hat{B}_N - B_1 = (X'_0 X_0)^{-1} X'_0 (Y - XM) = (X'_1 N X_1)^{-1} X'_0 N^{1/2} W_N$$

and  $C_N(\hat{B}_N - B_1)A = C_N (X'_1 N X_1)^{-1} X'_1 N^{1/2} W_N A = D_N W_N A$ ,

where

$$D_N = C_N (X'_1 N X_1)^{-1} X'_1 N^{1/2}.$$

The c.f.  $\phi_N(T_1)$  of  $C_N(\hat{B}_N - B_1)A$  is

$$\begin{aligned} \phi_N(T_1) &= E[\exp(i \text{Tr}(T'_1 D_N W_N A))] = \zeta_N(D'_N T_1 A) = \\ &= \exp(-0.5 \text{Tr}(D'_N T_1 A \Sigma A T'_1 D_N)) + f_N(D'_N T_1 A), \end{aligned}$$

from (9),

$$= \exp(-0.5 \text{Tr}(T_1 T'_1)) + f_N(D'_N T_1 A),$$

since, from (11)  $A \Sigma A = I_p$  and  $D_N D'_N = C_N (X'_1 N X_1)^{-1} C_N = I_r$ .

To show that  $\lim_{N \rightarrow \infty} f_N(D'_N T_1 A) = 0$  for fixed  $T_1$ , note first that

$$\text{Tr}((D'_N T_1 A)' (D'_N T_1 A)) = \text{Tr}(T'_1 T_1 \Sigma^{-1}) \quad \forall N.$$

The result then follows from (9) by choosing

$$A = \left\{ T; \text{Tr}(T' T) \leq \text{Tr}(T'_1 T_1 \Sigma^{-1}) \right\}.$$

Of more interest in practice is the result obtained by replacing  $\Sigma$  by  $S$  (or  $S_0$ , which is readily seen to converge in probability to  $\Sigma$  when  $H$  is true). If we write  $A = \phi_2(\Sigma)$  and define

$$A_N = \phi_2(S), \quad (12)$$

it follows immediately from Theorem 2 that  $A_N \xrightarrow{D} A$  and that

$$C_N(\hat{B}_N - B_1)A_N = (C_N(\hat{B}_N - B_1)A)A^{-1}A_N \xrightarrow{D} Z_1 A^{-1}A = Z_1,$$

which proves the following result.

**Theorem 8.** When  $H$  is true

$$\hat{B}_N \xrightarrow{L} N(B_1, S \otimes (X'_1 N X_1)^{-1}),$$

in the sense that

$$C_N (\hat{B}_N - B_1) A_N \xrightarrow{D} N(0, I_{pr}),$$

where  $C_N$  and  $A_N$  are defined in (11) and (12).

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#### STRESZCZENIE

W pracy prezentuje się uogólnienie zbieżności według rozkładu na wielowskaźnikowe tablice wektorów losowych. Rozważania te wykorzystuje się w analizie zbieżności według rozkładu statystyki  $T_0^2$ -Hotellinga i innych statystyk (w przypadku rozkładu różnego od normalnego) wykorzystywanych w MANOVA.

#### РЕЗЮМЕ

В работе представляются обобщение сходимости по распределению на мультииндексные таблицы случайных векторов. Эти исследования используются в анализе сходимости по распределению статистики  $T_0^2$ -Хотеллинга и других статистик (в случае распределения разного от нормального) использованных в MANOVA.

